

# *Predictability of South China Sea summer monsoon onset*

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1 **Predictability of South China Sea Summer Monsoon onset**

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11 **ABSTRACT**

12 Predicting monsoon onset is crucial for agriculture and socioeconomic planning in  
13 countries where millions rely on the timely arrival of monsoon rains for their livelihoods.  
14 In this study we demonstrate useful skill in predicting year to year variations in South China  
15 Sea summer monsoon onset at up to 3 months lead time using the GloSea5 seasonal  
16 forecasting system. The main source of predictability comes from skilful prediction of  
17 Pacific sea surface temperatures associated with El Niño and La Niña. The South China  
18 Sea summer monsoon onset is a known indicator of the broadscale seasonal transition that  
19 represents the first stage of the onset of the Asian summer monsoon as a whole. Subsequent  
20 development of rainfall across East Asia is influenced by sub-seasonal variability and  
21 synoptic events that reduce predictability, but interannual variability in the broadscale

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22 monsoon onset for East Asian summer monsoon still provides potentially useful  
23 information for users about possible delays or early occurrence of the onset of rainfall over  
24 East Asia.

25 **Key words:** SCSSM, South China Sea Summer Monsoon, EASM, East Asian Summer  
26 Monsoon

27

## 28 **1. Introduction**

29 The broadscale East Asian Summer Monsoon (EASM) onset occurs in two stages  
30 (Wang et al, 2004; 2009): The first stage is a seasonal transition that occurs over the South  
31 China Sea (SCS) and is characterised by an abrupt but sustained reversal of the lower  
32 tropospheric zonal winds from easterlies to westerlies. Several studies have considered the  
33 SCS Summer Monsoon (SCSSM) onset as the precursor for the EASM development (Tao  
34 and Chen, 1987; Lau and Yang, 1997), with the formation and progression of the mei-yu  
35 rainband forming the second salient phase (Wang et al., 2004). Predicting monsoon onset  
36 is crucial for agriculture and socioeconomic planning in countries where millions rely on  
37 the timely arrival of monsoon rains for their livelihoods.

38 Interannual variability in the seasonal transition that constitutes the broadscale  
39 monsoon onset has been shown to be related to thermal conditions over the Tibetan Plateau  
40 (Wu et al., 2012), El Niño-Southern Oscillation (ENSO) effects (Zhou and Chan, 2006; Hu  
41 et al., 2014; Xie et al., 2015; Zhu and Li, 2017), regional air-sea interactions (He and Wu,  
42 2013) and intraseasonal oscillations (ISO; Li et al., 2013; Wu 2010; Zhu and He, 2013;  
43 Shao et al., 2014; Wang et al., 2017). He et al. (2017) carried out a comprehensive analysis

44 of the SCSSM onset in individual years between 1997 and 2014 and showed that the years  
45 can be divided into “normal”, “intermittent” and “delayed” onset years based on the  
46 development of local circulations, thermodynamic conditions and rainfall patterns  
47 following the seasonal transition. He et al. (2017) found that eight out of the 18 years they  
48 analyzed exhibited intermittent rainfall onset (such that the seasonal dynamical transition  
49 is not closely followed by the establishment of monsoon rains and maximum SCS surface  
50 temperatures, with a delay caused by an active ISO or northern cold air entering the SCS),  
51 and suggested that this reduces the potential predictability of local rainfall onset even if the  
52 seasonal dynamical transition may be predictable. Wang et al. (2017) described the effects  
53 of the tropical ISO on early, normal and late SCSSM onsets observed over 34 years. They  
54 confirmed work from previous studies which showed that, before each onset, the SCS is  
55 controlled by the dry phase of the ISO (Shao et al., 2014), and the SCS is warmed to  
56 precondition the onset, while after each onset, the SCS is cooled by the wet phase of the  
57 ISO (Wu, 2010). However, Wang et al. (2017) showed that the transition process is found  
58 to be related to different ISO evolutions over the Indian Ocean for the three types of onsets.

59 Even in non-intermittent onset years, the progression of rainfall onset over East Asia  
60 is rarely smooth. After an initial burst of rainfall over the SCS, the rain band rapidly  
61 advances northward before stagnating over the Yangtze and Huai River valleys in the mei-  
62 yu front (baiu in Japan). The mei-yu rainband exhibits large intra-seasonal and interannual  
63 variability and has been the subject of extensive literature (see Ding and Chan (2005) for a  
64 review). Its onset is associated with a northward shift of the Northwest Pacific Subtropical  
65 High axis to about 25°N and the migration of the upper level westerly jet over Eurasia to  
66 the north of the Tibetan Plateau (Sampe and Xie, 2010; Luo 2013). Li et al. (2018) showed

67 that the anticyclone in the upper troposphere over South Asia in April has a significant  
68 relationship with the mei-yu onset dates, such that a stronger South Asian anticyclone in  
69 April is followed by earlier onset dates of the mei-yu.

70 Despite the complexity associated with these multiple drivers, interannual variability  
71 in the seasonal transition that constitutes the broadscale monsoon onset for the Asian  
72 summer monsoon as a whole still provides useful information for forecasters about possible  
73 delays or early occurrence of the onset of rainfall over East Asia. One of the most-used  
74 indices for determining SCSSM onset is that proposed by Wang et al. (2004). This index  
75 identifies the first pentad after 25th April in which the zonal wind at 850 hPa over the  
76 southern part of the South China Sea ( $5^{\circ}$ - $15^{\circ}$ N,  $110^{\circ}$ - $120^{\circ}$ E) shifts from a mean easterly to  
77 a mean westerly. Wang et al. (2004) demonstrate that this index is highly indicative of the  
78 seasonal transition of the large-scale circulation. They showed that the onset variations  
79 determined using this index matched the broadscale onset determined by the principal  
80 component of the first empirical orthogonal function (EOF) of the low level winds over  
81 East Asia and the Western North Pacific. They argued that this simple index avoids the  
82 additional complications of the intraseasonal variability that is included in EOF analysis.

83 An alternative definition for SCSSM onset was proposed by Gao et al. (2001), and is  
84 used in operational extended-range forecasting by the Chinese Meteorological  
85 Administration (D. Zhang, personal communication). This includes an additional criterion  
86 of a sustained increase of equivalent potential temperature at 850 hPa above 340K over the  
87 SCS region  $10^{\circ}$ - $20^{\circ}$ N,  $110^{\circ}$ - $120^{\circ}$ E concurrent with the establishment of westerly winds  
88 over the same region. The increase in equivalent potential temperature is considered to  
89 indicate sea surface warming, monsoonal transport of moisture into the region and the

90 potential for increased convective activity (Gao et al., 2001; Luo et al., 2013; Li et al.,  
91 2013). The region specified by Gao et al. (2001) is further north than that for the Wang et  
92 al. (2004) index and includes the northern SCS.

93 In this paper, we investigate the prediction skill of the SCSSM onset on seasonal  
94 timescales in the operational hindcast set of the GloSea5-GC2 seasonal forecasting system.  
95 Section 2 outlines the data and methods used in our study; section 3 shows the analysis of  
96 predictability of the two onset indicators, including tests of the robustness of the seasonal  
97 forecast skill. Discussion and conclusions on the usefulness of the seasonal forecast skill  
98 of the broadscale monsoon onset using these SCSSM onset indicators are included in  
99 Section 4.

## 100 **2. Data and methods**

101 Daily and pentad timeseries of 850 hPa zonal winds ( $U_{850}$ ), air temperature ( $T_{850}$ ) and  
102 specific humidity ( $q_{850}$ ) from the 23-year set of hindcasts (1993-2015) made with the  
103 GloSea5-GC2 operational long-range forecast system (MacLachlan et al, 2015; Williams  
104 et al., 2015) are taken from four start dates (17<sup>th</sup>, 25<sup>th</sup> March, 1<sup>st</sup>, 9<sup>th</sup> April). This represents  
105 a >1-month lead-time for the average SCSSM onset date of mid-May. The standard  
106 operational hindcast set includes 7 members per start date. In order to investigate the  
107 robustness of our results, and the dependence on ensemble size, we make use of an  
108 additional hindcast ensemble, using the same model configuration and also with 7 members  
109 per start date (except for 17<sup>th</sup> March, for which there are only 3 members). Further, to  
110 investigate changes with lead-time, we repeat the analysis for a 56-member ensemble of  
111 start dates 25<sup>th</sup> March, 1<sup>st</sup>, 9<sup>th</sup>, 17<sup>th</sup> April, and for 28-member ensembles generated using  
112 the four start dates (1<sup>st</sup>, 9<sup>th</sup>, 17<sup>th</sup>, 25<sup>th</sup>) of January, February and March respectively.

113 Data representative of observations are taken from ERA-Interim reanalyses (Dee et  
114 al., 2011) for the same years. Equivalent potential temperature is calculated from the  
115 temperature and humidity fields at 850 hPa using the formula in Bolton (1980). Sea surface  
116 temperatures for March (used in section 3.3) are taken from the HadISST1.1 dataset  
117 (Rayner et al., 2005).

118 SCSSM onset is determined using the criteria established by Wang et al. (2004) and  
119 Gao et al. (2001). According to Wang et al. (2004), the onset date is the first pentad after  
120 25th April (i.e. pentad 24 onwards) when the area-averaged  $U_{850}$  over the southern SCS  
121 ( $5^{\circ}$ - $15^{\circ}$ N,  $110^{\circ}$ - $120^{\circ}$ E, denoted  $U_{SCS}$ ) is (a)  $> 0 \text{ m s}^{-1}$ ; (b) in the subsequent four pentads  
122 (including the onset pentad)  $U_{SCS}$  must be positive in at least three pentads, and (c) the  
123 accumulated 4-pentad mean  $U_{SCS} > 1 \text{ m s}^{-1}$ . Wang et al (2004), He and Zhu (2015) and Zhu  
124 and Li (2017) have compared the SCSSM onset pentads between different reanalyses  
125 (including the National Centers for Environment Prediction (NCEP) reanalyses versions I  
126 (Kalnay et al., 1996) and II (Kanamitsu et al., 2002) as well as ERA-Interim) and show  
127 reasonable correlations between them (generally  $>0.8$ ).

128 Gao et al. (2001) suggested an onset criterion based on the area-averaged 850 hPa  
129 pentad equivalent potential temperature ( $\theta_e$ ) and  $U_{850}$  over the region  $10^{\circ}$ - $20^{\circ}$ N,  $110^{\circ}$ -  
130  $120^{\circ}$ E, with the onset date being the first pentad when  $\theta_e > 340\text{K}^{\dagger}$  and the  $U_{850} > 0.0 \text{ m s}^{-1}$   
131 stably (persists for at least three pentads followed by a break of no more than 2 pentads, or  
132 for two pentads followed by a break of no more than one pentad). It should be noted that  
133 the region of consideration for this index is slightly further north than that considered by  
134 Wang et al. (2004).

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<sup>†</sup> Originally specified as 335K by Gao et al. (2001) but revised to 340K by Ding and He (2006).



### 135 3. Results

#### 136 *3.1 Prediction skill of SCSSM onset using the Wang et al. (2004) criterion*

137 Figure 1 shows the SCSSM onset pentads identified using the Wang et al. (2004)  
138 criterion for each forecast member with start dates 17<sup>th</sup>, 25<sup>th</sup> March, 1<sup>st</sup>, 9<sup>th</sup> April in each  
139 year, with the ensemble mean pentad and that identified in the reanalyses. The average  
140 interannual standard deviation of onset dates from individual ensemble members is 2.2  
141 pentads, which compares reasonably well with that of the reanalyses (2.6 pentads), and  
142 there is a statistically significant (at the 0.75% level, for a one-tailed t-test) correlation of  
143 0.5 between the interannual variations of the ensemble mean dates and those from the  
144 reanalyses, indicating significant predictability. The hindcasts also predict the mean onset  
145 pentad to match that of the reanalyses, i.e. pentad 28 (16<sup>th</sup> - 20<sup>th</sup> May).

146 Luo and Lin (2017) suggest that a more objective measure of the SCSSM onset can be  
147 determined using a daily cumulative  $U_{SCS}$  and specifying the onset as where this time series  
148 changes from decreasing to increasing (indicating that the flow is becoming predominantly  
149 westerly). Wang et al. (2004) also checked their SCSSM onset dates against a cumulative  
150  $U_{SCS}$  criterion, DU, which compares the accumulated  $U_{SCS}$  in the 3 days prior to and after  
151 the onset. They showed that although their onset criteria do not explicitly require an abrupt  
152 change in westerly speed across the onset pentad, the resultant onset pentads were  
153 coincident with such a change. We find that including the additional criterion of  $DU > 7 \text{ m}$   
154  $\text{s}^{-1}$  makes very little difference to our results (not shown).

155 We have carried out the same analysis for four start dates (1<sup>st</sup>, 9<sup>th</sup>, 17<sup>th</sup>, 25<sup>th</sup>) in January,  
156 February and March taken from the standard operational hindcast ensemble of 7 members  
157 per start date, and also for a 56 member combined ensemble using start dates of 25<sup>th</sup> March,

158 1<sup>st</sup>, 9<sup>th</sup> and 17<sup>th</sup> April (see Table 1). The correlation coefficient increases with decreasing  
159 lead-time, becoming statistically significant at the 1.5% level (for a one-tailed t-test) from  
160 February start dates onwards. Thus, there is significant skill in the SCSSM onset prediction  
161 using the Wang et al. (2004) index at nearly 3 months lead-time over this hindcast period.

### 162 ***3.2 Predictability of SCSSM onset using the Gao et al. (2001) criterion***

163 Figure 2 shows the SCSSM dates identified using the Gao et al. (2001) criterion in  
164 each year by each of the 52 ensemble members with start dates 17<sup>th</sup>, 25<sup>th</sup> March, 1<sup>st</sup>, 9<sup>th</sup>  
165 April, with the ensemble mean pentad and that identified in the reanalyses. In contrast with  
166 the findings using the Wang et al. (2004)  $U_{SCS}$  index, we find low skill in onset prediction  
167 using the Gao et al. (2001) index at >1 month lead time. Table 1 shows that the correlation  
168 increases slightly if the lead-time is reduced to ~1 month, but remains barely statistically  
169 insignificant at the 6% level (using a one-tailed t-test).

170 The difference in prediction skill between the two methods of determining SCSSM  
171 onset may be in part related to the region used for the Gao et al. (2001) index; Wang et al.  
172 (2004) commented that “*the northern SCS is open to the invasion of a cold front from the*  
173 *north. The westerly flow occurring before the onset is located north of the subtropical ridge*  
174 *and is not of tropical origin.*” They state, therefore, that the northern part of the SCS should  
175 be excluded when defining the tropical monsoon burst over the SCS. He et al. (2017) also  
176 commented on the influence of northern cold air entering this region of the SCS  
177 contributing to ambiguous or intermittent onset. They highlighted the case of 2009, where  
178 the strong westerly flow established in mid-April was interrupted by easterlies propagating  
179 from the northern SCS for several days in early May. Other examples of years where this  
180 occurred were given in He et al. (2017, their Figures 1, 2) and include 2007, 2009, 2011.

181 Additionally, although He et al (2017) did not identify 2004 as an intermittent onset year,  
182 the  $U_{850}$  averaged over the Gao et al. (2001) SCS box fluctuates between easterly and  
183 westerly during May, making the onset ambiguous when the Gao et al. (2001) index is  
184 used. He et al. (2017, their Figure 1) shows that this is related to variability of the winds in  
185 the northern part of the SCS. In contrast, the  $U_{850}$  winds over the southern part of the SCS  
186 (as covered by the Wang et al. (2004) box) do not fluctuate to the same extent. Chan et al.  
187 (2000) showed that, in 1998, incursion of cold air into the northern SCS promoted release  
188 of convectively available potential energy which helped to trigger the onset earlier than  
189 may have been expected given the ENSO conditions. Liu et al. (2002) further linked the  
190 cold air incursion to a Rossby wave train triggered over the Bay of Bengal.

191 The additional influence of variability from the subtropics in the northern SCS, which,  
192 like the ISO, is unpredictable on seasonal timescales, is likely to be a contributing factor in  
193 the reduced seasonal prediction skill for SCSSM onset using the Gao et al. (2001) criteria.  
194 In recognition of this, forecasters at CMA release their SCSSM onset forecasts using the  
195 Gao et al. (2001) criteria only on the extended range (11-30 day) timescale (D. Zhang,  
196 personal communication, 30<sup>th</sup> March 2018), on which models have been shown in previous  
197 work to have skill for predicting intraseasonal variability (e.g. Lim et al., 2018; Lee et al.  
198 2015).

### 199 ***3.3 Drivers of SCSSM onset predictability using Wang et al. (2004) index***

200 Several studies have shown that ENSO is one of the main drivers of large-scale  
201 interannual variability in the Asian monsoon region (e.g. Zhou and Chan, 2007; Luo et al.,  
202 2016). Westerly (easterly) equatorial wind anomalies associated with El Niño (La Niña)  
203 and a weaker (stronger) Walker circulation are typically associated with negative (positive)

204 sea surface temperature (SST) anomalies over the SCS and a delayed (advanced) seasonal  
205 transition (He et al., 2017). This relationship is not symmetrical, however: He et al. (2017)  
206 suggest that both intraseasonal oscillations (ISO) and changes in west–east thermal  
207 contrasts across the Indian Ocean and western Pacific can influence the timing of onset in  
208 La Nina years. Hardiman et al. (2018) found a similar asymmetry in the relationship  
209 between seasonal mean Yangtze River rainfall and ENSO in observations and hindcasts.

210 We also show on Figure 1 the observed March Niño3.4 sea surface temperature (SST)  
211 anomaly timeseries from HadISST1.1 (yellow line). The correlation coefficient between  
212 the ensemble mean SCSSM onset pentad timeseries derived using the Wang et al. (2004)  
213 index and the Niño3.4 SST timeseries is 0.9, indicating that the predictable component of  
214 the hindcast SCSSM onset is driven mainly by ENSO, which itself is highly predictable on  
215 this timescale in GloSea5 (MacLachlan et al., 2015; Scaife et al., 2014). The correlation  
216 between observed estimates of SCSSM onset and the observed March Niño3.4 SST is  
217 rather lower (0.41), indicating the influence of other drivers of SCSSM onset variability  
218 that may not be predictable, particularly the ISO (e.g. Shao et al., 2014; Wang et al., 2017),  
219 which is itself subject to inter-annual variations relating to large-scale modes such as the  
220 Pacific-Japan teleconnection (Li et al., 2014). The skill of the ensemble (0.5) is therefore  
221 marginally higher than using predicted ENSO conditions alone to predict monsoon onset,  
222 though both are skilful.

223 Figure 3(a) provides additional insight by showing the correlation between the  
224 ensemble mean SCSSM onset dates for the 23 years from the hindcast and observed global  
225 monthly mean SSTs in March over the same period. This illustrates that the predictable  
226 part of the SCSSM onset from the hindcast is strongly correlated with an ENSO-like pattern

227 of Pacific SSTs, consistent with the findings of Zhu and Li (2017). There is also a strong  
228 positive correlation with SSTs in the equatorial Indian Ocean, again indicating that warmer  
229 SSTs are associated with later SCSSM onset dates. For the observed onset dates derived  
230 from ERA-interim (Fig. 3(b)), the correlations with SST are far smaller, due to the presence  
231 of additional factors in the observations that are not predicted by the ensemble mean. The  
232 average correlations between the SSTs and 1000 pseudo-timeseries of SCSSM onset  
233 created by randomly choosing an individual ensemble member hindcast for each year (Fig.  
234 3(c)) are naturally smaller than with the ensemble mean timeseries, but not as low as those  
235 in observations (Fig. 3(b)). This suggests that some of the sub-seasonal variations (e.g.  
236 intraseasonal oscillations) that affect SCSSM onset in reality may not be sufficiently well  
237 represented by the model to capture such influences, even at the relatively high horizontal  
238 resolution used by GloSea5 (N216; about 60 km at 50°N). This is consistent with findings  
239 of Fang et al. (2016), who showed that while several aspects of the boreal summer ISO  
240 were improved in the Met Office Unified Model at this resolution, difficulty remained in  
241 realistic representation of the variance and propagation characteristics.

#### 242 ***3.4 Robustness of SCSSM wind onset predictability to ensemble size***

243 To assess the influence of ensemble size on the prediction skill using the Wang et al.  
244 (2004) index, we randomly sample small ensembles of between 1 and 51 members from  
245 the 52 members in our combined ensembles with start dates between 17<sup>th</sup> March and 9<sup>th</sup>  
246 April, and re-calculate the correlation between the ensemble-mean timeseries and that from  
247 the observations for different numbers of ensemble members. Figure 4 indicates that, for  
248 this measure of monsoon onset, the prediction skill (black line) rises quickly with ensemble  
249 size, reaching a mean value of 0.5 for a 28-member ensemble (which is the size of the

250 standard operational hindcast set), and is robust (correlation coefficients averaged over all  
251 ensemble-mean timeseries are statistically significant at the 1% level for a one-tailed test)  
252 for around 10 ensemble members or more. This is a reflection of the strong and predictable  
253 influence of ENSO on wider tropical rainfall (Kumar et al., 2013; Scaife et al, 2017) and  
254 here on the SCSSM onset dates in the hindcast: in most of the summers following strong  
255 El Niño/La Niña years (e.g. 1998, 1999, 2000, 2001, 2005, 2008, 2010) the spread among  
256 ensemble members is small and several members identify the same onset pentad (see  
257 Figure 1), thereby constraining the values selected by random sampling of the ensemble  
258 for those years.

259 Several authors (e.g. Scaife et al., 2014; Eade et al., 2014; Dunstone et al., 2016) have  
260 demonstrated that the model's North Atlantic Oscillation is less predictable than that  
261 observed, so that a large number of ensemble members is required for good prediction skill.  
262 This was confirmed by repeatedly randomly selecting a single member to be the truth and  
263 using the ensemble mean of the remaining members to predict that member. In contrast,  
264 the dashed line on Figure 4 indicates that the model's SCSSM onset dates are more  
265 predictable than those from reanalyses, *i.e.* that the model is *over-confident* in its  
266 predictions, as is often found for tropical rainfall (Weisheimer and Palmer, 2014). This  
267 again illustrates the dominant role of ENSO in providing the predictability in the model,  
268 while the observed onset dates are also influenced by intraseasonal variations that are  
269 unpredictable on the seasonal timescale.

#### 270 **4. Conclusions**

271 SCSSM onset, as determined by the Wang et al. (2004)  $U_{850}$  wind index, is skilfully  
272 predicted in GloSea5 at up to 3 months lead time, particularly during active ENSO years.

273 Since the SCSSM onset signifies the start of the broadscale EASM, its skilful prediction is  
274 important for forecasters as an indicator of the possible characteristics of the season to  
275 come. This complements the skill previously demonstrated for predicting seasonal mean  
276 precipitation in the Yangtze River region (Li et al., 2016). The prediction skill for SCSSM  
277 onset using this index is robust even with only around 10 ensemble members, consistent  
278 with skill in prediction of rainfall in the deep tropics (e.g. Scaife et al., 2017). The skill is  
279 largely related to ENSO SSTs which have been shown to be highly predictable in the  
280 GloSea5 seasonal forecasting system.

281 In contrast, the Gao et al. (2001) SCSSM onset index, which includes an increase of  
282  $\theta_e$  in the SCS region as a measure of thermodynamic onset alongside the change to westerly  
283 winds, shows little predictability on seasonal timescales. We speculate that this is partly  
284 due to the region used by Gao et al. (2001), as this includes the northern SCS which can be  
285 influenced by incursions of cold air from the north. This additional influence is, like the  
286 ISO, inherently unpredictable on the seasonal timescale, and thus its inclusion through the  
287 northward extension of the box used for the Gao et al. (2001) index compared with that of  
288 Wang et al. (2004) is, in our view, a contributing factor in the reduced seasonal prediction  
289 skill. However, we propose that a seasonal forecast of the broadscale transition using the  
290 Wang et al. (2004) index would provide some useful early information for forecasters, and  
291 their guidance could later be refined, using other measures such as the Gao et al. (2001)  
292 index, with medium-range forecasts that may capture the influence of intraseasonal  
293 variations at shorter lead-times.

294 He and Zhu (2015) investigated the correlations between the SCSSM onset (as  
295 determined by the Wang et al. (2004) criteria) and the subsequent EASM rainfall from May

296 to September in observations/reanalyses. They suggested that, in contrast with the  
297 traditional view that a later onset date would be associated with a lower than normal total  
298 seasonal rainfall amount, the region from the lower Yangtze River to Korea and southern  
299 Japan shows a positive correlation between the SCSSM onset date and the seasonal mean  
300 rainfall, i.e. early SCSSM onset tends to be followed by lower than normal seasonal mean  
301 rainfall further north. He and Zhu (2015) associate this relationship with a persistent  
302 Western North Pacific anticyclonic/cyclonic anomaly accompanied by decaying El  
303 Niño/La Niña conditions in boreal spring to summer (Wu et al., 2010; Stuecker et al., 2013;  
304 Hardiman et al, 2017). This suggests that skilful predictions of SCSSM onset could provide  
305 an indication of the seasonal mean rainfall in parts of the EASM region.

306 To our knowledge, this is the first time that skill in predicting the broadscale transition  
307 associated with the SCSSM onset on seasonal timescales in an operational dynamical  
308 forecasting system has been demonstrated. We encourage other centres to investigate this  
309 in their operational forecasting systems. While it is recognised that the onset and  
310 progression of the SCSSM and EASM systems is complex and may be influenced by other  
311 factors such as synoptic events, intraseasonal variability and regional air-sea interactions  
312 with little or no predictability on the seasonal timescale, the ability to provide skilful  
313 predictions of whether the broadscale seasonal transition is likely to be early, late or normal  
314 provides useful, early information for local forecasters, particularly when combined with  
315 other predictions, such as the Yangtze River basin rainfall, which have also been shown to  
316 be skilful (Li et al., 2016) and are now provided in real time to CMA (Bett et al., 2018).

317



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### FIGURE CAPTIONS

485 **Figure 1:** Predictability of the SCSSM wind onset: onset pentads derived using the method  
 486 proposed by Wang et al. (2004) from the GloSea5 ensemble predictions initialized on  
 487 17th, 25th March, 1st, 9th April (green dots represent individual members of the 52-  
 488 member ensemble, with the size of the dot scaled by the number of members predicting  
 489 the same onset pentad) and their ensemble mean (green line) compared with the  
 490 equivalent onset pentads derived from ERA-Interim (black line). The yellow line  
 491 shows the Niño3.4 SST anomaly in March for each year taken from the HadISST1.1  
 492 dataset. Pearson correlation coefficients are given in the legend:  $r(\text{ens,obs})$  represents  
 493 the correlation between the GloSea5 ensemble mean and ERA-Interim;  $r(\text{ens,sst})$   
 494 represents the correlation between the GloSea5 ensemble mean SCS onset pentads and  
 495 the observed March Niño3.4 SST anomaly;  $r(\text{obs,sst})$  represents the correlation  
 496 between the ERA-Interim SCS onset pentads and the observed March Niño3.4 SST  
 497 anomaly.

498 **Figure 2:** As Fig. 1 but for SCSSM thermodynamic onset as determined by a sustained  
 499 increased of  $\theta_{e,SCS}$  above 340K accompanied by the establishment of westerly winds



500 over the region 10-20N 110-120E, as proposed by Gao et al. (2001) (with the threshold  
501 modified by Ding and He, 2006).

502 **Figure 3:** Correlation coefficients between SCSSM onset pentad derived using the Wang  
503 et al. (2004) index and observed March average sea surface temperatures from  
504 HadISST1.1 for the period 1993-2015, using: (a) ensemble mean onset dates from the  
505 hindcast; (b) onset dates from ERA-Interim, (c) 10,000 pseudo-timeseries of onset  
506 dates created by randomly selecting an individual ensemble member from each year;  
507 panel shows average over all correlations. Contours and darker shades indicate  
508 correlations significant at the 1% ( $r=0.48$ ) and 3% ( $r=0.40$ ) levels respectively, for a  
509 one-tailed t-test.

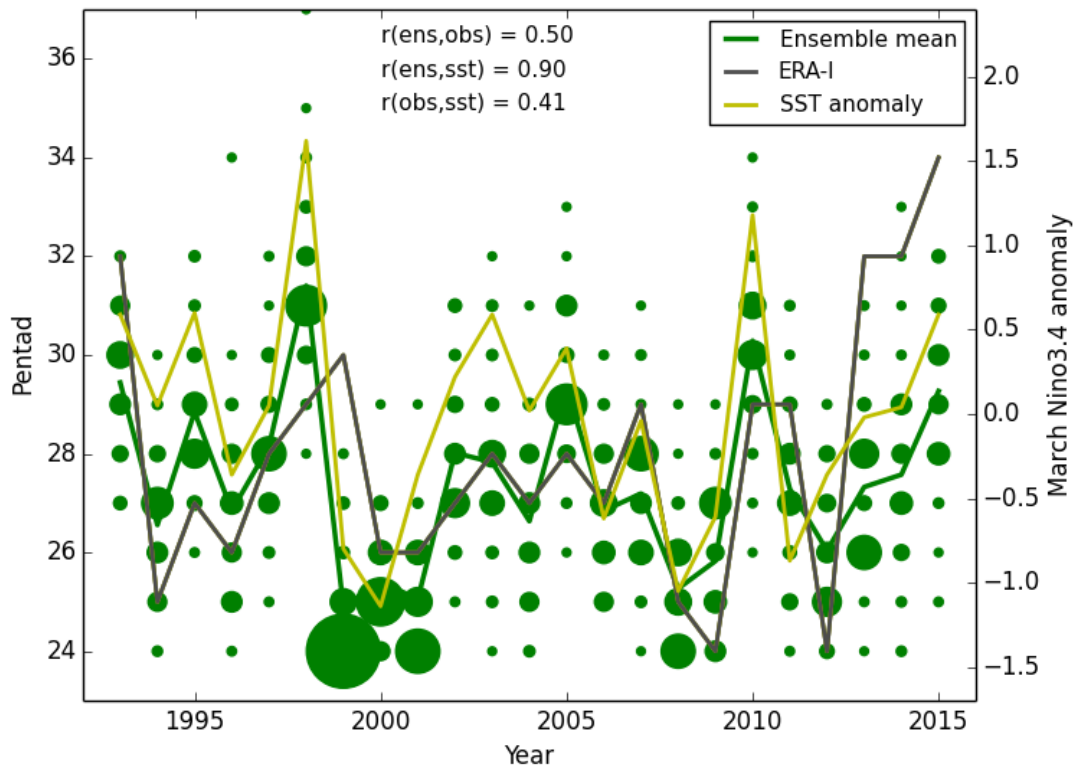
510 **Figure 4:** Effect of ensemble size on the skill of SCSSM onset predictions using the Wang  
511 et al. (2004) index (solid line), denoted  $r(\text{ens,obs})$ , and the signal to noise ratio  
512 (correlation of ensemble mean timeseries with a pseudo-timeseries created by  
513 randomly selecting a single model ensemble member for each year, dashed line),  
514 denoted  $r(\text{ens,mod})$ . In both cases, for each choice of ensemble size, up to 10,000  
515 ensemble-mean timeseries are generated by randomly selecting the chosen number of  
516 ensemble member onset dates (independently and without replacement) from the 52  
517 onset dates diagnosed in each year in the combined ensemble and averaging over the  
518 chosen number of ensemble members. Dot-dashed lines indicate the values of  $r$  that  
519 are significant at the 1% and 0.1% levels for a one-tailed t-test.

520

521 **Table 1.** Pearson correlation coefficients between ensemble mean SCSSM onset dates  
 522 from GloSea5 and those from ERA-Interim, using the definitions of Wang et al. (2004)  
 523 and Gao et al. (2001), for different hindcast start dates. Note that the earliest observed  
 524 SCSSM onset date is pentad 25 (1<sup>st</sup> -5<sup>th</sup> May) and the mean onset date is pentad 28 (16<sup>th</sup>-  
 525 20<sup>th</sup> May). Where just the month is shown, start dates are 1<sup>st</sup>, 9<sup>th</sup>, 17<sup>th</sup>, and 25<sup>th</sup> of the month.  
 526 Correlation coefficients statistically significant (for a 23 year hindcast period) at <1.5%  
 527 level for a 1-tailed test are in *italics* and those significant at <1% level for a 1-tailed test  
 528 are in **bold**.

	Ensemble start dates				
	January	February	March	17 <sup>th</sup> , 25 <sup>th</sup> March, 1 <sup>st</sup> , 9 <sup>th</sup> April	25 <sup>th</sup> March, 1 <sup>st</sup> , 9 <sup>th</sup> , 17 <sup>th</sup> April
Wang et al. (2004)	0.28	<i>0.46</i>	<i>0.45</i>	<b>0.50</b>	<b>0.53</b>
Gao et al. (2001)	-	-	-	0.27	0.30

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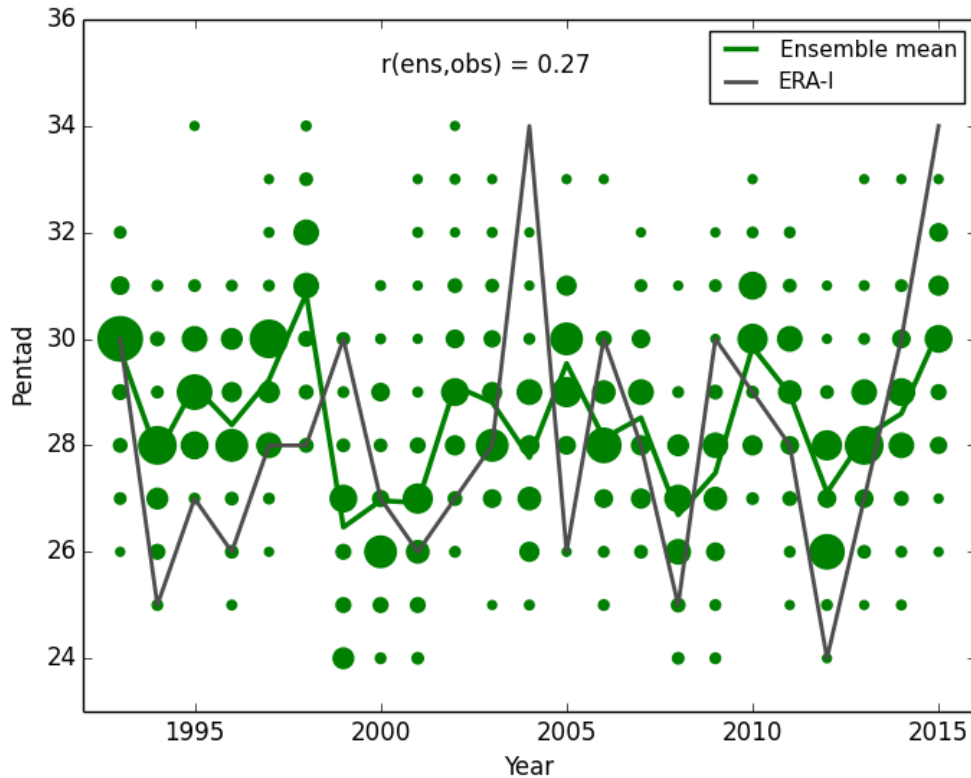
530

531 **Fig. 1:** Predictability of the SCSSM wind onset: onset pentads derived using the method  
 532 proposed by Wang et al. (2004) from the GloSea5 ensemble predictions initialized on 17<sup>th</sup>,  
 533 25<sup>th</sup> March, 1<sup>st</sup>, 9<sup>th</sup> April (green dots represent individual members of the 52-member  
 534 ensemble, with the size of the dot scaled by the number of members predicting the same  
 535 onset pentad) and their ensemble mean (green line) compared with the equivalent onset  
 536 pentads derived from ERA-Interim (black line). The yellow line shows the Niño3.4 SST  
 537 anomaly in March for each year taken from the HadISST1.1 dataset. Pearson correlation  
 538 coefficients are given in the legend:  $r(\text{ens,obs})$  represents the correlation between the  
 539 GloSea5 ensemble mean and ERA-Interim;  $r(\text{ens,sst})$  represents the correlation between  
 540 the GloSea5 ensemble mean SCS onset pentads and the observed March Niño3.4 SST

541 anomaly;  $r(\text{obs}, \text{sst})$  represents the correlation between the ERA-Interim SCS onset pentads  
 542 and the observed March Niño3.4 SST anomaly.

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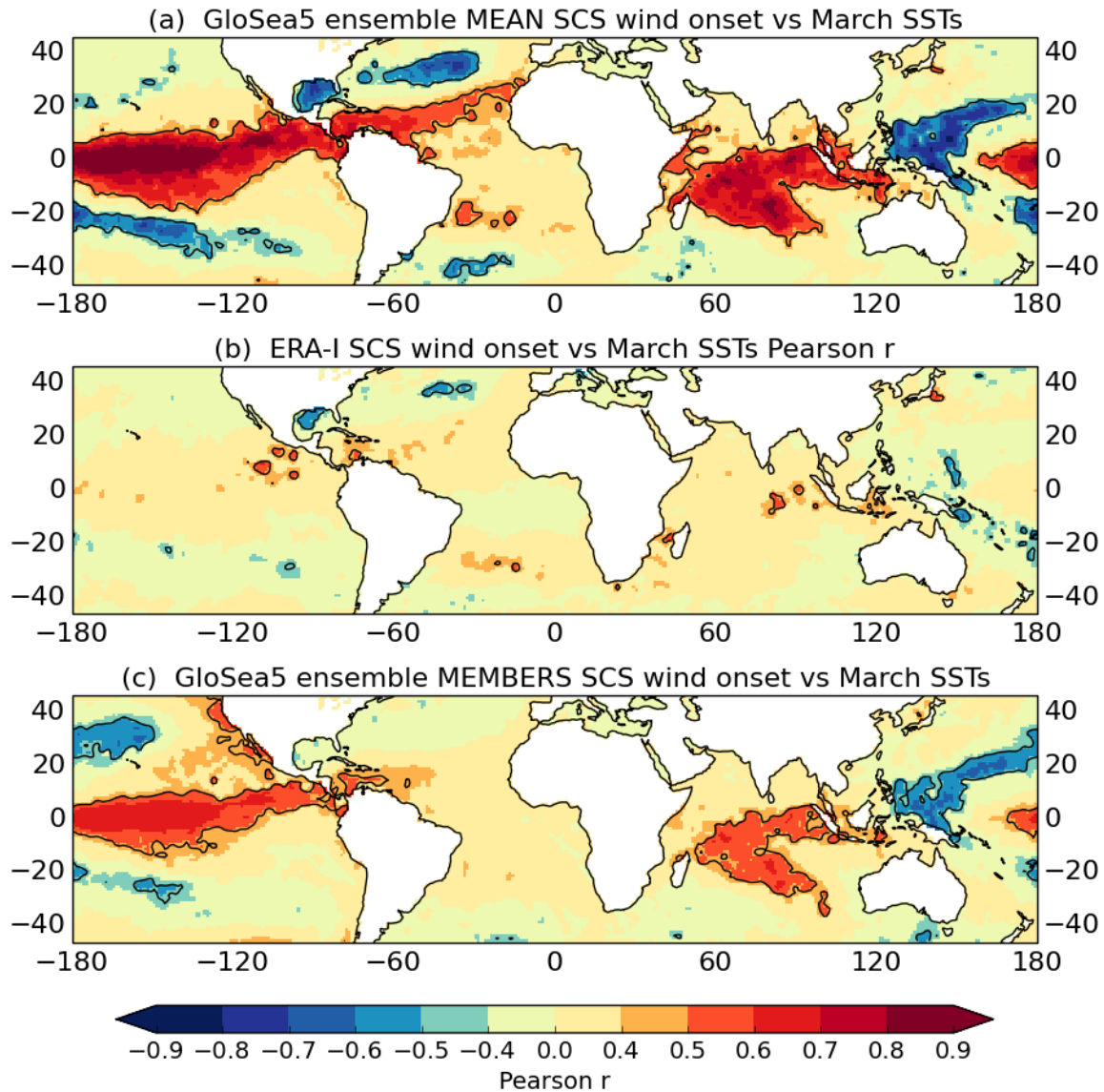
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546 **Fig. 2:** As Fig. 1 but for SCSSM thermodynamic onset as determined by a sustained  
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 548 the region 10-20N 110-120E, as proposed by Gao et al. (2001) (with the threshold modified  
 549 by Ding and He, 2006).

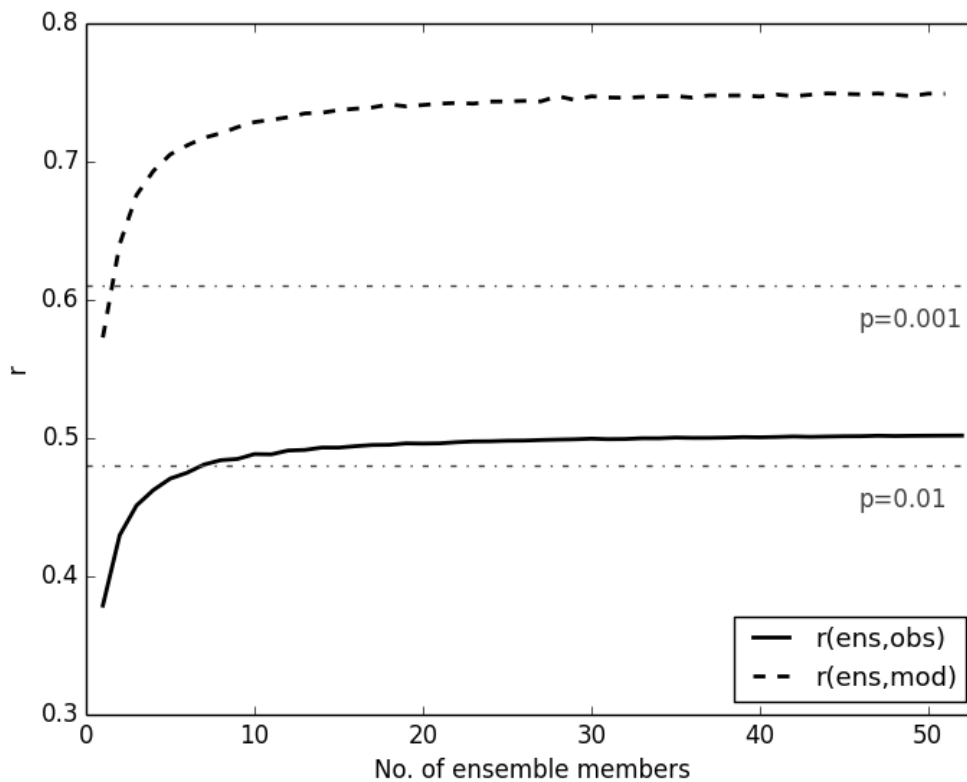
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552 **Fig. 3:** Correlation coefficients between SCSSM onset pentad derived using the Wang et  
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 554 for the period 1993-2015, using: (a) ensemble mean onset dates from the hindcast; (b) onset  
 555 dates from ERA-Interim, (c) 10,000 pseudo-timeseries of onset dates created by randomly  
 556 selecting an individual ensemble member from each year; panel shows average over all  
 557 correlations. Contours and darker shades indicate correlations significant at the 1%  
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560

561 **Fig. 4:** Effect of ensemble size on the skill of SCSSM onset predictions using the Wang  
 562 et al. (2004) index (solid line), denoted  $r(\text{ens,obs})$ , and the signal to noise ratio  
 563 (correlation of ensemble mean timeseries with a pseudo-timeseries created by randomly  
 564 selecting a single model ensemble member for each year, dashed line), denoted  
 565  $r(\text{ens,mod})$ . In both cases, for each choice of ensemble size, up to 10,000 ensemble-mean  
 566 timeseries are generated by randomly selecting the chosen number of ensemble member  
 567 onset dates (independently and without replacement) from the 52 onset dates diagnosed  
 568 in each year in the combined ensemble and averaging over the chosen number of  
 569 ensemble members. Dot-dashed lines indicate the values of  $r$  that are significant at the  
 570 1% and 0.1% levels for a one-tailed t-test.